

# CSE 591 - FALL 03.

Chitta Baral

Department of Computer Science and Engineering  
Arizona State University  
Tempe, AZ 85287-5406 USA  
chitta@asu.edu

<http://www.public.asu.edu/~cbaral/cse571-f99/>

October 12, 2003

# PROBABILITY, BAYES NETS AND CAUSALITY

## Basic Concepts in probability theory

- 3 basic axioms of probability calculus in the Bayesian formalism
  - $0 \leq P(A) \leq 1$
  - $P(\text{Sure proposition}) = 1$
  - $P(A \vee B) = P(A) + P(B)$ , if  $A$  and  $B$  are mutually exclusive.
    - \*  $P(A) = P(A, B) + P(A, \neg B)$   
( $P(A, B)$  is short for  $P(A \wedge B)$ )
    - \* If  $B_i, i = 1, 2, \dots, n$  is a set of exhaustive and mutually exclusive propositions (called a partition or a variable), then

$$P(A) = \sum_i P(A, B_i)$$

- Basic expression in Bayesian formalism
  - Conditional probabilities of the form  $P(A|B)$

- means: belief in  $A$  under the assumption that  $B$  is known with absolute certainty.
- $P(A|B) = P(A)$  –  $A$  and  $B$  are independent.
- $P(A|B, C) = P(A|C)$  –  $A$  and  $B$  are conditionally independent given  $C$ .
- Dawid's notation:  $(A \perp\!\!\!\perp B|C)$
- Bayesian philosophers see the conditional relationship as more basic than that of joint events.

$$P(A \wedge B) = P(A|B)P(B)$$

## Bayesian Networks

- Goal:
  - to provide convenient means of expressing substantive assumptions
  - to facilitate economical representations of joint probability functions
  - to facilitate efficient inferences from observations
- Idea: Directed acyclic graphs is used to represent causal or temporal relationship
- Basic decomposition scheme
  - $P(A \wedge B) = P(A|B)P(B)$
  - $P(x_1, x_2, x_3) = P(x_1 \wedge x_2 \wedge x_3) = P(x_1|x_2, x_3)P(x_2 \wedge x_3) = P(x_1|x_2, x_3)P(x_2|x_3)P(x_3)$

– In general,

$$P(x_1, \dots, x_n) = \prod_j P(x_j | x_1, \dots, x_{j-1})$$

\* Let  $PA_j \subseteq \{x_1, \dots, x_{j-1}\}$ , such that  $x_j$  is independent of  $\{x_1, \dots, x_{j-1}\} \setminus PA_j$  once we know the value of  $PA_j$ .

\* We can then write

$$P(x_j | x_1, \dots, x_{j-1}) = P(x_j | pa_j)$$

\* If  $PA_j$  is a minimal set of predecessors of  $X_j$  that renders  $X_j$  independent of all its other predecessors, then  $PA_j$  is said to be **Markovian parents** of  $X_j$ .

- Markov factorization: If a probability function  $P$  admits the factorization

$$P(x_1, \dots, x_n) = \prod_j P(x_j | parents_j)$$

relative to a DAG  $G$ , we say  $G$  represents  $P$ , that  $G$  and  $P$  are compatible, or  $P$  is Markov relative to  $G$ .

## Inference with Bayesian Networks

- Prediction and abduction
  - $x$  – a set of observations
  - $y$  – a set of variables deemed important for prediction or diagnosis
  - Need to compute  $P(y|x)$ .
  -

$$P(y|x) = \frac{p(y, x)}{p(x)} = \frac{\sum_s P(y, x, s)}{\sum_{y,s} P(y, x, s)}$$

- An example:
  - The Network
    - \*  $P(\textit{tampering}) = 0.02$ ;  $P(\textit{fire}) = 0.01$
    - \* Directed Edges:  $(\textit{tampering}, \textit{alarm})$ ,  $(\textit{fire}, \textit{alarm})$ ,  
 $(\textit{fire}, \textit{smoke})$ ,  $(\textit{alarm}, \textit{leaving})$ ,  $(\textit{leaving}, \textit{report})$

\* local probability distributions:

$$P(\text{alarm}|\text{fire}, \text{tampering}) = 0.5;$$

$$P(\text{alarm}|\text{fire}, \neg \text{tampering}) = 0.99;$$

$$P(\text{alarm}|\neg \text{fire}, \text{tampering}) = 0.85;$$

$$P(\text{alarm}|\neg \text{fire}, \neg \text{tampering}) = 0.0001.$$

$$P(\text{smoke}|\text{fire}) = 0.9; P(\text{smoke}, \neg \text{fire}) = 0.01.$$

$$P(\text{leaving}|\text{alarm}) = 0.88; P(\text{leaving}|\neg \text{alarm}) = 0.001.$$

$$P(\text{report}|\text{leaving}) = 0.75; P(\text{report}|\neg \text{leaving}) = 0.01.$$

– Different kinds of inferences

\* Diagnostic inferences:  $P(\text{fire}|\text{report})$

\* Causal inferences (prediction):  $P(\text{leaving}|\text{tampering})$

\* Intercausal inferences:  $P(\text{fire}|\text{alarm}, \text{tampering})$

\* Mixed inferences:  $P(\text{alarm}|\text{report}, \text{fire})$

– An illustration:

$$\begin{aligned} &P(\text{tampering}|\text{report}, \text{smoke}) \\ &= \frac{P(\text{tampering}, \text{report}, \text{smoke})}{P(\text{report}, \text{smoke})} \end{aligned}$$



$$= \frac{\Sigma_{leaving, alarm, fire} P(tampering=T, report=T, smoke=T, leaving, alarm, fire)}{\Sigma_{tampering, leaving, alarm, fire} P(report=T, smoke=T, tampering, leaving, alarm, fire)}$$

\* Let us compute the denominator  $D$  first.

$$\begin{aligned} & \Sigma_{tampering, leaving, alarm, fire} P(tampering) P(fire) \\ & P(smoke = T|fire) P(alarm|tampering, fire) \\ & P(leaving|alarm) P(report = T|leaving) \\ & = \Sigma_{tampering, leaving, alarm} P(tampering) P(leaving|alarm) \\ & P(report = T|leaving) \Sigma_{fire} P(fire) P(smoke = T|fire) \\ & P(alarm|tampering, fire) \end{aligned}$$

\* Let  $f_1(alarm, tampering) = \Sigma_{fire} P(fire) P(smoke = T|fire) P(alarm|tampering, fire)$

$$\begin{aligned} & \text{Now let us compute } f_1(alarm = T, tampering = T) \\ & = \Sigma_{fire} P(fire) P(smoke = T|fire) \\ & P(alarm = T|tampering = T, fire) \\ & = P(fire = T) P(smoke = T|fire = T) \\ & P(alarm = T|tampering = T, fire = T) + \\ & P(fire = F) P(smoke = T|fire = F) \\ & P(alarm = T|tampering = T, fire = F) \end{aligned}$$

$$= 0.01 \times 0.9 \times 0.5 + 0.99 \times 0.01 \times 0.85$$

Similarly, we can also compute  $f_1(\text{alarm} = T, \text{tampering} = F)$ ,  $f_1(\text{alarm} = F, \text{tampering} = T)$  and  $f_1(\text{alarm} = F, \text{tampering} = F)$ .

\* We can now write the denominator as:

$$\begin{aligned} & \Sigma_{\text{tampering}, \text{leaving}, \text{alarm}} P(\text{tampering}) P(\text{leaving} | \text{alarm}) \\ & P(\text{report} = T | \text{leaving}) f_1(\text{alarm}, \text{tampering}) \\ & = \Sigma_{\text{tampering}, \text{leaving}} P(\text{tampering}) P(\text{report} = T | \text{leaving}) \Sigma_{\text{alarm}} \\ & P(\text{leaving} | \text{alarm}) f_1(\text{alarm}, \text{tampering}) \end{aligned}$$

Let us denote  $\Sigma_{\text{alarm}} P(\text{leaving} | \text{alarm}) f_1(\text{alarm}, \text{tampering})$  by  $f_2(\text{leaving}, \text{tampering})$ . We can compute it as we compute  $f_1$

\* The denominator can now be written as:

$$\begin{aligned} & = \Sigma_{\text{tampering}, \text{leaving}} P(\text{tampering}) P(\text{report} = T | \text{leaving}) \\ & f_2(\text{leaving}, \text{tampering}) \\ & = \Sigma_{\text{tampering}} P(\text{tampering}) \Sigma_{\text{leaving}} P(\text{report} = T | \text{leaving}) \\ & f_2(\text{leaving}, \text{tampering}) \end{aligned}$$

Let us denote  $\Sigma_{\text{leaving}} P(\text{report} = T | \text{leaving})$

$f_2(\textit{leaving}, \textit{tampering})$  by  $f_3(\textit{tampering})$  and compute it like the other  $f_i$ s.

\* The denominator can now be written as:

$$\sum_{\textit{tampering}} P(\textit{tampering}) f_3(\textit{tampering})$$

- Main Issues and challenges

- Computing the conditional probabilities efficiently
- Inference in general networks is NP-hard
- Many efficient algorithms are defined for particular kind of networks (say for trees).
  - \* Algorithm based on message passing architecture for trees.
  - \* Join-tree propagation
  - \* Cutset conditioning
  - \* Hybrid combinations of the above two
  - \* Approximation methods: stochastic simulation.

## Causal Bayesian Networks

- Motivation

- A joint distributions tells us how probable events are and how probabilities would change with subsequent observations.
- A causal model also tells us how these probabilities would change as a result of external interventions.

Such a change can not be deduced from a join distribution even if fully specified.

- Importance

- Difference between **observing** the alarm is on, and **turning** the alarm on.
- $P(\textit{fire}|\textit{alarm}) > 0.01$ .

But  $P(\textit{fire}|\textit{do}(\textit{alarm} = T)) = P(\textit{fire}) = 0.01$

- Causal networks can predict the effect of actions. (Simple joint distributions can not.)
- Stability and autonomy
  - Autonomy: It is possible to change one parent child relationship in the network without changing the others.
  - Stability: One can predict the effect of external interventions with minimum of extra information.
  - Autonomy and intervention: Instead of specifying a new probability function for each of the many possible interventions, we specify merely the immediate changes implied by the intervention. Because of autonomy, the change is local.
- Definition: Causal Bayesian network

Let  $P(v)$  be a probability distribution on a set  $V$  of variables, and let  $P_x(v)$  denote the distribution resulting from the intervention  $do(X = x)$  which sets any subset  $X$  of variables to constants  $x$ .

Denote by  $P^*$  the set of all interventional distributions  $P_x(v)$ ,  $X \subseteq V$ ,

including  $P(v)$  which represents no intervention. A DAG  $G$  is said to be a **causal Bayesian network** compatible with  $P^*$  iff the following three conditions hold for every  $P_x \in P^*$ .

1.  $P_x(v)$  is Markov relative to  $G$ .
2.  $P_x(v_i) = 1$ , for all  $V_i \in X$ , whenever  $v_i$  is consistent with  $X = x$ .
3.  $P_x(v_i|pa_i) = P(v_i|pa_i)$  for all  $V_i \notin X$ , whenever  $pa_i$  is consistent with  $X = x$ .

- Properties:

- for all  $v$  consistent with  $x$ :

$$P_x(v) = \prod_{\{i|V_i \notin X\}} P(v_i|pa_i)$$

- For all  $i$ ,  $P(v_i|pa_i) = P_{pa_i}(v_i)$

(The above ensures, conditional probabilities with respect to parents, corresponds to causal effects.)

- For all  $i$ , and for every subset  $S$  of variables disjoint of  $\{V_i, PA_i\}$  we have:  $P_{pa_i, S}(v_i) = P_{pa_i}(v_i)$   
(Expresses invariance of causality)
- Causal relationship is more stable than probabilistic relationships.
  - Causal relationship remains unaltered as long as no change has taken place in the environment, even when our knowledge about the environment undergoes change.
    - \*  $(season, sprinkler), (season, rain), (sprinkler, wet), (rain, wet), (wet, slippery)$ .
    - \*  $S_1$  – Turning the sprinkler on would not affect rain
    - \*  $S_2$  – The state of the sprinkler is independent of the state of the rain.
    - \*  $S_2$  changes from false to true when we learn what season it is.
    - \* Given that we know the season,  $S_2$  changes from true to false once we observe that the pavement is wet.

- \*  $S_1$  remains true regardless of what we learn or know about the season or the pavement.
- \* Falling barometer predicts rain, does not explain it.



## Functional Causal Models

- Two views of non-determinism
  - Laplace's (1814) conception of natural phenomena:  
Nature's laws are deterministic, and randomness surfaces merely due to our ignorance of the underlying boundary condition.
  - Modern (quantum mechanical) conception of physics:  
All relationships are inherently stochastic.
- Why Pearl's book uses Laplace's conception of causality
  - besides the fact that it is used in genetics, econometrics and social sciences
  - It is more general.
    - \* Every stochastic model can be emulated by many functional relationships (with stochastic inputs), but not the other way round;

- \* Functional relationships can only be approximated as a limiting case, using stochastic models.
- Laplacian conception is more in tune with human intuition.
- Certain important concepts can only be defined in Laplacian framework (i.e., they can not be defined in terms of purely stochastic models.)
  - \* the probability that event  $B$  occurred *due to* event  $A$ .
  - \* the probability that event  $B$  would have been different if it were not for event  $A$   
(they are called counterfactuals)
- (Functional) causal model:  
A causal model is a triple  $M = \langle U, V, F \rangle$  where
  - $U$  is a set of background (or exogenous, or error ) variables, that are determined by factors outside the model.
  - $V$  is a set  $\{V_1, \dots, V_n\}$  of variables, that are determined by the variables in  $U \cup V$ .

- $F$  is a set of functions  $\{f_1, \dots, f_n\}$  giving rise to a set of structural equations of the form:  $x_i = f_i(pa_i, u_i)$ ,  $i = 1, \dots, n$
- Types of queries that can be answered using functional causal models
  - **Prediction:** Would the pavement be slippery if we *find* the sprinkler off?
  - **Interventions:** Would the pavement be slippery if we *make sure* that the sprinkler is off?
  - **Counterfactuals:** Would the pavement be slippery *had* the sprinkler been off, given that the pavement is in fact not slippery and the sprinkler is on?
- Prediction using Markovian causal models:
  - Causal diagram: A graph obtained by having edges from each member of  $PA_i$  to  $X_i$ .
  - If the causal diagram is acyclic then the corresponding model is called semi-Markovian.

- \* the values of  $X$  variables will be uniquely determined by the  $U$  variables.
- \* The joint distribution  $P(x_1, \dots, x_n)$  is determined uniquely by the distribution  $P(u)$  of the error variables.
- If in addition the error terms are mutually independent, the model is called *Markovian*.
- Theorem (Pearl and Verma): Every Markovian causal model  $M$  induces a distribution  $P(x_1, \dots, x_n)$  that satisfies the Markov condition relative to the causal diagram  $G$  associated with  $M$ , that is each variable  $X_i$  is independent on all its non-descendants, given its parents  $PA_i$  in  $G$ .
- Theorem (Drudgel and Simon): For every Bayesian network  $G$  characterized by a distribution  $P$ , there exists a function model that generates a distribution identical to  $P$ .
- Advantages of doing prediction using causal-functional specification over the probabilistic specification
  - \* When organizing knowledge using Markov causal models reliable

assertions about conditional independence can be made without assessing numerical probabilities. (They come later when writing what  $f$  exactly is and what the  $P(u_i)$ 's are.)

- \* Functional specification is often more meaningful, natural and yields a smaller number of parameters.
- \* Judgemental assumptions of conditional independence of observable quantities are simplified, and made more reliable, when cast directly as judgments about the presence or absence of *unobserved* common causes. (Instead of judging whether each variable is independent of all its nondescendants, given its parents, we need to judge whether the parent set contains *all* relevant immediate causes, namely whether two omitted factors (say  $U_i$  and  $U_j$ ) share a common cause.
- \* When some conditions in the environment undergo change, it is simpler to reassess (judgmentally) or reestimate (statistically) the model parameters knowing that the change is local, affecting just a few parameters, than reestimating the whole model from

scratch.

- Interventions and causal effects in functional models.

- Submodels of causal models:

Let  $M$  be a causal model,  $X$  be a set of variables in  $V$ , and  $x$  be a particular realization of  $X$ . A submodel  $M_x$  of  $M$  is the causal model  $M_x = \langle U, V, F_x \rangle$ , where

$$F_x = \{f_i : V_i \notin X\} \cup \{X = x\}.$$

- Effects of actions on a causal model: The effect of action  $do(X = x)$  on a causal model  $M$  is given by the submodel  $M_x$ .

- Effects of actions on other variables: The potential response (or value) of a variable  $Y$  in  $V$  after an action  $do(X = x)$  denoted by  $Y_x(u)$  is the solution for  $Y$  using the set of equations  $F_x$ .

- Advantages over stochastic models

- \* The analysis of interventions can be directly extended to cyclic models.

$$(demand = f(price, income, u_1); price = f'(demand, cost, u_2))$$

- \* Analysis of causal effects in non-Markovian models will be greatly simplified using functional models.

(Because: There are infinitely many conditional probabilities  $P(x|pa_i)$ , but only finite number of functions  $x_i = f_i(pa_i, u_i)$ , among discrete variables  $X_i$  and  $PA_i$ .)

- **Counterfactuals**

- Why we can not use causal Bayes nets.
  - \* Counterfactuals involve dealing with both actions and observations. (Effect of a drug on a patient with certain symptoms.)
  - \* The observations alter the conditional probabilities.
- An example illustrating the inadequacy of using causal Bayes nets.
  - \*  $X$  denotes a treatment.
  - \*  $Y = 0$  means recovery and  $Y = 1$  means death.
  - \* Q: A certain patient Joe, took the treatment and died. Our question is whether Joe's death occurred *due* to the treatment.

I.e., What is the probability that Joe (or any patient for that matter) , who died under treatment ( $x = 1, y = 1$ ) would have recovered ( $y = 0$ ) had he not been treated ( $x = 0$ ).

- \* An extreme case: 50% of the patients recover and 50% die in both the treatment and the control groups. (assume sample size to be infinite.) I.e.  $P(y|x) = \frac{1}{2}$ .
- \* Bayes net 0: edge-less, with  $P(y, x) = 0.25$ , for all  $x$  and  $y$
- \* Functional model 1:  $x = u_1, y = u_2$ , with  $P(u_1 = 1) = P(u_2 = 1) = \frac{1}{2}$ .
- \* Functional model 2:  $x = u_1, y = xu_2 + (1 - x)(1 - u_2)$ , with  $P(u_1 = 1) = P(u_2 = 1) = \frac{1}{2}$ .
- \* Both functional model 1 and 2 correspond to the same joint probability  $P(y, x) = 0.25$ , for all  $x$  and  $y$ . But will give different answers.



\* Answering Q using model 1 and model 2

$y$	$u_2$	$x$	$P_{model1}(y u_2, x)$	$P_{model2}(y u_2, x)$
0	0	0	0.25	0
0	0	1	0.25	0.25
<u>0</u>	<u>1</u>	<u>0</u>	<u>0</u>	<u>0.25</u>
0	1	1	0	0
1	0	0	0	0.25
<u>1</u>	<u>0</u>	<u>1</u>	0	0
1	1	0	0.25	0
<u>1</u>	<u>1</u>	<u>1</u>	0.25	0.25

\* Using model 1 the answer to Q would be 0.

Intuitively: the treatment has no effect. 50% die and 50% recover.

\* Using model 2 the answer to Q would be 1.

Intuitively, the treatment kills 50% of the people and cures the other 50%.

- Answering counter-factual queries using functional models.
  - Counterfactual: Let  $Y$  be a variable in  $V$  in the causal model  $M = \langle U, V, F \rangle$ . The counterfactual sentence “The value that  $Y$  would have obtained, had  $X$  been  $x$ ” is interpreted as denoting the potential response  $Y_x(u)$ .
  - Probabilistic causal model: Is a pair  $\langle M, P(u) \rangle$ , where  $M$  is a causal model and  $P(u)$  is a probability function defined over the domain of  $U$ .
    - \*  $P(y) = P(Y = y) = \sum_{\{u | Y(u)=y\}} P(u)$
    - \*  $P(Y_x = y) = \sum_{\{u | Y_x(u)=y\}} P(u)$
    - \*  $P(Y_x = y, X = x') = \sum_{\{u | Y_x(u)=y \& X(u)=x'\}} P(u)$
    - \*  $P(Y_x = y, Y_{x'} = y') = \sum_{\{u | Y_x(u)=y \& Y_{x'}(u)=y'\}} P(u)$
  - One purpose of counter-factuals: We want to show that the event  $X = x$  was *the cause* of the event  $Y = y$ .
  - So we ask the question: What is the probability that  $Y$  would not be equal to  $y$  had  $X$  not been equal to  $x$ ?

- To answer the above we need to evaluate  $P(Y_{x'} = y' | X = x, Y = y)$
- Given  $M$ , a three step procedure to evaluate the conditional probability  $P(B_A | e)$  of a counterfactual sentence “If it were A then B,”, given evidence  $e$ .
  - (  $e$  is  $X = x$  and  $Y = y$ .  $A$  is  $X \neq x$ .)
  - \* Abduction: Update  $P(u)$  by the evidence  $e$ , to obtain  $P(u|e)$ .  
(explain the past ( $U$ ) in light of the current evidence  $e$ .)
  - \* Action: Modify  $M$  by the action  $do(A)$  to obtain  $M_A$ .  
(minimally bend the course of history, to comply with the hypothetical condition  $X \neq x$ )
  - \* Prediction: Use the modified model  $\langle M_A, P(u|e) \rangle$  to compute the probability of  $B$ .  
(predicting the future ( $Y$ ) on the basis of the above 2 steps.)

## Evaluating Counter-factuals: an example

- The Causal relationship in a 2-man firing squad:
  - Nodes
    - \*  $U$  : Court orders the execution.
    - \*  $C$  : Captain gives a signal.
    - \*  $A$  : Rifleman- $A$  shoots.
    - \*  $B$  : Rifleman- $B$  shoots.
    - \*  $D$  : Prisoner dies.
  - Edges:  $(U, C), (C, A), (C, B), (A, D), (B, D)$ .
- Logical structural equations
  - $C \Leftrightarrow U$
  - $A \Leftrightarrow C$
  - $B \Leftrightarrow C$

$$- D \Leftrightarrow A \vee B$$

- Questions that we want to answer:
  - (prediction) : If the rifleman did not shoot, the prisoner would be alive.
  - (abduction) : If the prisoner is alive, then the captain did not signal.
  - (transduction) : If rifleman- $A$  shot, then  $B$  shot as well.
  - (action) : If the captain gave no signal and rifleman- $A$  decides to shoot, the prisoner will die and  $B$  will not shoot.
  - (counter-factual) : If the prisoner is dead, then even if  $A$  were not to have shot, the prisoner would still be dead.
- Probabilistic analysis: a modification of the story
  - There is a probability  $P(u = 1) = p$  that the court has ordered the execution.
  - Rifleman- $A$  has a probability  $q$  of pulling the trigger out of nervousness. ( $w = 1$ )

- Rifleman-A's nervousness is independent of  $U$ .
- We wish to compute the probability that the prisoner would be alive if  $A$  were not to have shot, given that the prisoner is in fact dead.
- The solution steps:
  - \* (abduction) :  $P(u, w|D)$
  - \* (action) :
  - \* (prediction) :